STATISTICAL DESIGN OF EXPERIMENT

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A TOOL FOR MINERAL ENGINEERS

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INTRODUCTION:

The classical methods of one-factor-at-a-time experiment(keeping all other remaining factors constant) requires a large number of trial involving time, energy, and money, yet the effect of interaction between various factors are not brought out clearly. In an effort to find optimum conditions with less number of experiments and to secure amount of quantitative information about the system, statistical or factorial design of experiments has been devised, A regression equation is developed:

Yi=(x1,x2,...xi...xn)

 $Yi = \beta o + \Sigma \beta i \times i + \Sigma \beta i i \times i^2 + \Sigma \beta i j \times i \times j$

Yi=bo+ bixi+Σbiixi²+Σbijxixj

where bo,bi,bij,bii are the best estimates(method of lowest squares) of regression coefficients &o,&i,&ij,&ii.

Efficiency of experimental investigation can be significantly improved by use of experimental design and data analysis based on statistical principles.

It is a mathematical method of drawing valid conclusion from a series of tests made in a predetermined pattern. It is a standard tool for the industrial experimenter as a chemical balance is for laboratory experimenter. * Very helpful for preliminary investigation/screening experiment of systems with several independent variables. It effectively determines which factors are important so that these can then be more thoroughly investigated.

* Reduction in number of tests to be conducted(i.e. simultaneous rather than classical mwthod of varrying one by one)

Experiments are well organised.

* Regression coefficients in conjunctions with existing knowledge or prior information can give better insight to the physical/ physicochemical phenomena occuring in the system and thus aid in quantifying the effects and interactions of various dactors. * Improvement in productivity and reliability of results. Thus leading to the shortest path in optimization of the variable The process, suitable for industrial research(more information at less cost and time)

FACTORIAL EXPERIMENTS:

By virtue of the literature already available or experience gained on similar systems, one decides about the various factors the range of variation(+1 for upper level, -1 for lower level, '0'zero for base level) in coded form for each factor- the range chosen preferably being made narrow so that variation of the response variable is

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expected to be linear. While selecting the range of variation of the factor, first the base level is selected around which one wishes to vary the factor depending on prior information through literature or experience and then the interval or variation is select

yj=zj-zj/ Azj

 x_j =coded values, $z_j, z^\circ j$ are the actual or natural values for it the factor at any level, base level (0) Az_j=the range = (zmax-zmin)/2 , z_j^2 =(zmax+zmin)/2

 x_j = coded value, takes values like +1, -1, 1.414, -1.414 etc.

Once the regression equation Yi=(Yi) is found, it is decoded to give the equation in which the effect of actual factors are related to response variables. (output)

There are various methods of finding regression rquations.

(I) YATES METHOD:

Resides giving the values of regression coefficients b,bij, it gives the additional information of SS (Sum of Squares), m.s(mean squares)

* Write the response in standard order (1,a,b,ab,c,ac,bc,abc) in a column. '1' denotes all factors are -ve: 'a' means factor A is a upper level (+1), all other factors being at lower level(-1) and so on.

* Add these in pairs and then substract these in pairs(values of one ahead/succeeding from one behind proceeding in the table to form the next column.

* Continue similar operation one after another or number of columns' $_{equal}$ to the number of factors k. Let the values in the last column last column be l.

* Actual effect = t/2.

* Variance = $SS = 1^2/2!$, m.s = SS/df, SS = sum of squares; m.s = mean squares, df = degrees of freedom.

* Residuals = VT- Σ Vi,VT(Total variance)= Σ Y²-cF,cF(correction factor) correction factor = (Σ Y)²/2^K.

EY= the value in the last column against row marked '1'

EVi = Sumof variance due to all individuals and interactions

* F ratio = m.s /m.s residual, for a factor to be significant

its F should be greater than F

nl=df corresponding to larger m.s

n2=df " " smaller estimate m.s (Denominator)

FULL FACTORIAL DESIGN & METHODS OF STEEPEST ASCENT

BULK FLOTATION OF COHFLEX Cu Pb Zn ORES

FLOTATION CONDITIONS (CONSTANT)

Addition	Quantity kg/ton	Conditioning time(min)
Sodium Silicate	1.5	5
Na2S		2
CuSO4	13	5
Collector (varyin	g)	Э
Frother	.15	5

% Solids = 15, RPM = 1250

Factors:

		G	P	X	et dan ter an ar ter Mada	
		7 0	11	2.5		
	0	70	9	1.5		
	****	50	7	.5		
	Step	0 20 r	2	1.0		
4						
	X	Base	G	P	Х	Reco.
	1	- L	.1.	4		07 1
	2	. 4			4	.01 25
	2	+		- h -	4	-1111
	4	+		· _	4	64.63
	5	+	+	+		88.67
	6	+	+			88.38
	7	-f-	<u></u>	+		61.11
	8	+				55.67
	9	-1-	0	0	0	82.78
				1		

G=grind (%-200#) P=pH of slurry (ed. X =collector (Now.1P)

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YATES METHOD:

		1	2	3	2 (Col3) /8 =SS=ms,(df=1)	V.Ratio= ms/residual	Nodi(ir) V.ratio
ī	55.69	144.07	293.88	615.97	1373 6661	943 45	⊊A i 4
b ab	61.14 88.67	155.88	60.22 44.61	16.07	32.280612 23.770512	22.17 16.33	2.41
	6	15.97720).8* 723	8.08**			
C AC	64.63 91.25	32.69	5.74	28.21	99.475512 30.459012	68.32 20.92	6.81'.
bc abc	74.11 92.1	26.62	-5.16 -8.63	4.59 -3.47	2.6335125 1.5051125	1.81 1.0	
add ever	255.57	7 359.26	354.68 368.4	664.84			

Note: a=G, b=P, c=X significant

* All interactions and residuals pooled together to 58.34 and this is used as modified residual for calculation of modified F = 4.54 1,4,p=.05

		G	P	Х
	bj	13.13	2	3.52
-	Т	20.00	2	1.00
ē	steps	3		
ŀ	jT :	262.6	4	3.52
****	К	consta	ant	35.2
******	T'	7.5	. 1	. 1
1	New	7.5	. 1	. 1

limiting factor 0.1 kg/t, collector=x

				۲	latrix 11
				4	
	Test	G	Р	х	Recovery
10	(base)	70	9	1.5	82.98
		7.5	. 1	.1	
l 1		77.5	9.1	1.6	85.49
12		85.0	9.2	1.7	92.68
13		92.5	9.3	-1.8	92.23

Test 12 or 13 taken as optimum.

X test:

R=bo+b1G+b2P+b3X =76.99+13.13G+2P+3.52X

Test rur	G	P	Х	Robs	Rcalc	X=(Robs-Rcalc)/Rcalc
		4-	- -	92.1	95.64	.1310
	÷		+	91.25	91.64	.0016
		+	- 1 -	74.11	69.38	.3224
			+	64.63	65.38	.0086
	+	+		88.67	88.6	. 0000
	+			88.38	84.6	.0005
		+		61.14	62.34	.0231
	***		-	55.67	58.38	.1239
	***					.6111

<X 7, 95%=14.07

Thus it can be said with confidence that the result of matrix 1 and represented by above linear equation.

Ind method:

 $f_{actors} = x_0 = x_1 = x_2 + x_1 + x_2 = x_1 + x_2 + x_1 + x_1 + x_2 + x_1 + x_1 + x_2 + x_1 + x_1$

$$\begin{split} & \Sigma \times i j^2 = (j)^2 = N \\ & b j = \Sigma \times i j \chi i / \Sigma \times i j^2 = (j \chi) / (j)^2 = (j \chi) / N \\ & S^2 b j = Se^2 / \Sigma \times i j^2 = Se^2 / (j)^2 = Se^2 / N \\ & Se^2 = \Sigma (yi - y) / n - 1 = error mean square = \\ & replicate observations (n) at base level. \\ & t j = Eb j / Sb j] > t table at p = .05, df = n - 1 ... to be a significant \end{split}$$

coefficient in regression equation.

Subscriptj 0 1 2 ... 12 13 ... 123

(jy)or()

bj

tj

collect all significant coefficients to form a regression equation of "y(estimate)

^y=b0+b1x1+...+b12x1x2+.... Sr²=j=1ΣN(yj²-^y)²/(N-1)= residual variance, l= number of significant coefficients

yj= observed, ^y= calculated from regression equation.

F-Test: (test for adequacy of fit): The test is done for confirming reproducibility of results. The purpose is to show that there exists no significant variations from batch to batch at 5% confidence intervals due to personel or experimental errors.

F=Sr²/Se² \langle F | f1=N-1, f2=n-2, n= number of observations at base level. So $\hat{\gamma}$ adequately fits and in Then converted in $\hat{\gamma}$) to natural picele. Example:

Fa	acto	r	(Code ·		leve	15
		*****			upper	base	lower
					- - -	Ø	
7.Be	ento	nite	2 (4	8.707	8.0	7.29
7.Wa	ater		I	3	3.354	3.0	2.646
Te	est	No.	Xo	A	В	AB	Yg
	1		-+-	+	·}·	-4-	10.11
	2		-1-		+	-	6.31
	3		+	+			11.27
	4		+	-		÷	7.24
(j) N=4	/) 1	34.9	73	7.83	-2.09	.23	
AY)=	12-	2(-	-ve)	1			
	34.	93-3	2(6.0	\$1+7.2	4)= /.	8.5	
BY)==	34.	93-2	2(11)	27+7.	24)≕ -	2.09	
BA)=	34.	93-3	2(6.3	\$1+11.	27)= -	.23	
	j		Ø	A	В		AB
((jy)	34	4.93	7.83	-2.09		23
j=(j _γ	/)/4	{	3.73	1.96	-0.52		.06
			<u>v</u>	where	bj >	Abil	
bj							

Test No. Yg Α В 9.64 9.40 9.52 9.24 9.33 5 Ø Ø 5 6 7 8 9 Ø Ø

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Se² = $\frac{\partial^2}{\partial 1 + 10}$ = .02478 Sbj²= Se²/(j)²= Se²/N=.02478/4 Sbj= 0.07871 In order that bjj is to be significant,tj=bj/Sbj>t table p=.05 df=n-1=5-1=4 t table=2.78 at p=.05, df= n-1=5-1=4 bj>t table x Sbj | Δ bj|= t table x S bj = 2.75 × .07871 = 2188 $= \frac{bj}{2} > |\Delta bj|$

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CENTRAL COMPOSITE ROTATABLE DESIGN (C C R D)

A second order Orthogonal design is not rotatable and as a consequence the errors in y at experimental points on the response surface may be smaller than they are as determined from regression equation.

To make centre composite rotatable , $\alpha = 2 \frac{1}{4}$ or 2 $(\frac{1}{4})$ and n0 observations at base levels increased.

As before, in addition to 2 experiments, some more experiments are performed at $\pm \alpha$ points at base levels.

STEPS:

* Find (0y), (1y), (2y)....,(12y),....,(11y),... where '0' corresponds to xo where all x are +ve i.e +1

* $(\Sigma j j y) = (11y) + (22y)$

* Find a, a, a, a, a, a, rom Table- 1

* Find by , by , by , 5by , 5by , tab according to equa	quation	•
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Ь	5 ₆ ~	⁵ ь	$\pm \Delta b = \pm S_b$	
$b_{a} = a1(0y) - a_{2}\xi(jjy)$	$S_{to}^{2} = a1S_{2}^{2}$	560	± 660	
$b_j = a3(0y)$	S _{bj} =a35€	Sbj	± ∆bj	
b _{ij} = a4 (ijy) ₄≈	$S_{bij}^{2} = a4S_{bij}^{2}$	Bbij	生山好	
b)j = a5(jjy)+ a6 (Σjjy) -a2 (0y)	5 bjj = a75 e	9 ₆₎ ;	± △6jj	

+a5(jjy)+L, where $L = a6\xi(jy) - a2(0y)$, a constant.

b) = AN $[\Sigma \lambda^2 (k+2) (Oy) - 2\lambda C \Sigma (jjy)] = al (Oy) - aZ \Sigma (jjy)$ b) = CN (jy) = a3(jy) b) = $\lambda N C^2 (ijy) = a^2 (ijy)$ b) = AN $[C(k+2)\lambda -k) C^2 (jjy) + (1-\lambda) C^2\Sigma (jjy) - 2\lambda C (Oy)]$ = $a_5(jjy) + a_6(jjy) - a_{\lambda}(Oy)$ Where A= $[Z\lambda((k+2)\lambda - k)]^{-1}$, C= N/ Σx_{in}^2 $\lambda = k(n_i + n_{\lambda})/[(k+2)n_{\lambda}] = kN/[(k+2)n_{\lambda}]$ k = number of independent variables, n_i = number of centre points n_{λ} = number of peripheral points

N=n1+ n2

No of	No	of points in	n the	T	otal		L val	ue
k k	2k	Star(+/-L)	Centre	N ·		2k/4		
	fact	orial points 2k	s (base leve	1)				
2	. 4	4	5		13		1.414	
ä	8	6	6		20		1.682	
4	. 16	8	7		31		2.0	
5	32	10	8		50		2.378	
			3		and a state of the state of the state of the		1 21	
No.of variable k	" a1	a2	a3	a4	a5	121	a6	a
	nen ek skonomen som av de en ek forst forstaller en er er	Na anda pinao manana na kana kana kana kana kana kana					1999 - 1999 - 1997 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 - 1999 -	
2	0.2	0.1 0	.125	0.25	0.125	0,0	1875	0.1436
Э	0.16638	0.056791 0	0.073224	0.125	0.0625	0.00	06889	0.0695
4	0.142857	0.035714 0	0.041667	0.0625	0.03125	0.0	0372	0.035
5	0.0988	0.0191 0	.0231	0.0312	0.0156	0.00	014 (5.0172

Example:

Using the same example as was considered earlier for 2k experiment but in this $\pm \infty = 1.414$ was added, resulting in four more experiments at base levels were conducted.

LEVELS

+1 Ø -1 ±w

1.414 -1.414

Bent. x1 8.707 8.0 7.243 9 7 Moist.x2 3.354 3.0 2.645 3.5 2.5

+1, -1, 0, +1.414, -1.414 denote levelst(upper, lower, base, $\pm \infty$), $\infty = 1.4141$ at which experiments were conducted.

уg

No. xo x1 x2 x12

1 10.11 + + + + 2 + 6.31 + 3 + 11.27 34.93 4 7.24 + + 5 + 1.414 Ø 10.04 6 + -1.414 Ø 6.52 7 : + Ø 1.414 7.05 8 + Ø -1.414 8.82 9 + Ø Ø 9.64 10 + Ø Ø 7.4 114.53 11 + Ø Ø 9.52 79.6 12 + Ø Ø 9.24 13 Ø Ø 9.33 +

(ay) (1y) (2y) (12y)

		:	14	:	
су	iy	2у	11y	22y	12y
14.53	12.82	-4.55	68.07	66.73	23

Σ(iiy)=134**.8**

bo=a1(oy) - a22(iiy)= .2(114.53)- .1(134.8)= 9.426=bo

bi = a3(iy) = .125(iy) = .125(12.32) = 1.6025 = b1

.125(-4.55)= -.56875= b2

bij= b12=a4(ijy)=.25(-.23)=-.0575= b12

bii= $a5(iiy)+a6\Sigma(iiy)-a2(oy)$

.1251(iiy)+.0187(134.8)-.1(114.53)

.1251(68.07)-8.93224 = .416683=b11

.1251(66.73)-8.93224=-**.**584317=b22

t p=.05, y=2.776, 82 n-110 = s2g or s2e= .0247799

 s^2 bo=a1 s^2 y=.2

s²bi= a3 s²y=.125

s²bij=a4 s²y=.25	\mathbb{S}^2 \vee	√s²bj≕sbj	±ob=tsb
s²bijj= a7s²y=.1438	.0247799		
	· · · ·	.703987	, 19543
* 8 *	*	.055455	.154498
		.0787081	.2185
		.0596937	.1657
ál t			

bo	b1	b2_	b11	b22 -	b12	± bo	± bi	± bij
9.43	1.6	57	4167	584	057	.1954	.154	.218

neglecting b's <t bi, ie. only b12

yg= 9.43 + 1.60x1- .57x2- .42x21- .59x22

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Summarys

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e Contain

- * Statistical Design of Experiment is an efficient method with least number of experiments in a planned manner, one can find out the predominant effect of certain variables (xi) & for their interactions (xixj) which is not possible in classical experiment.
- * First build up Full/fractional factorial experiment to screen out important variables & interaction by t-test (significant coefficients) & F-test (adequacy of fit)
- * Use method of steepest ascent when first order coefficients are significant.
- * Otherwise continue some more tests at star points $\pm \alpha$ and use CCRD for 2nd order regression.

BIBLIDGRAPHY

1. Davies OL - Design and Analysis of Industrial Experiments, 101 Ltd, Longman group Ltd; NY, London. 2nd ed(1978)

- 2. Box & WilsonKP, J of Royal Society.Vol 13, 1, (1751)
- 3. Chew Victor, Experimental Design in Industry, John Wiley NY (1958)
- 4. Cochran WG & Cox GM, Experimental Design, Wiley NY (1950)
- 5. Kafarov, Cybernetic methods in analysis of Chemical Engineering Process, Mir Publishers (1976)
- 6. Fuersteral MC, Flotation vol 2, AM Gaudin Memorial Volume (AL Popular p427), NY: AIME 1976.

 Biles William & Swain JJ, Optimization and Industrial Experiment Wiley Inter Science (1980)

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